

Network Interference in Micro-Randomized Trials

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Nathan's Group Meeting

Joint work with

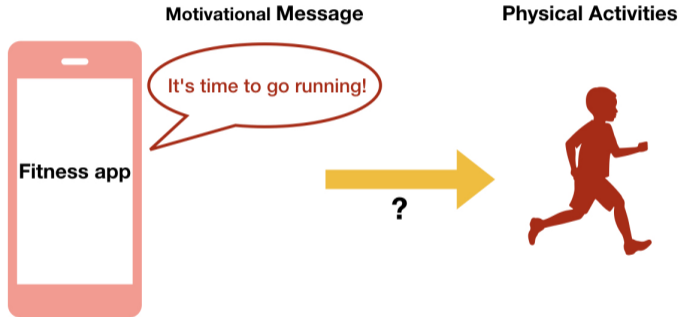


Stefan Wager

Li, S., & Wager, S. (2022). **Network Interference in Micro-Randomized Trials.** *arXiv preprint arXiv:2202.05356.*

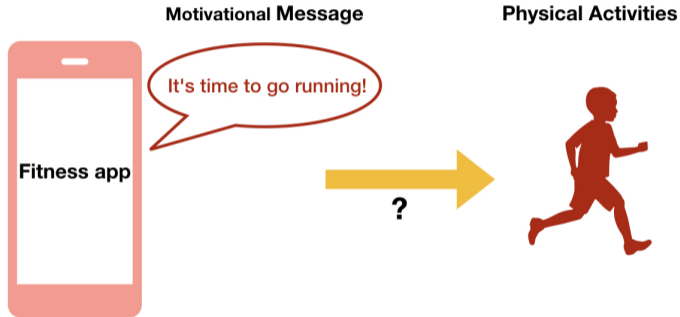
Motivation

- ▶ Effect of motivational message on level of physical activities



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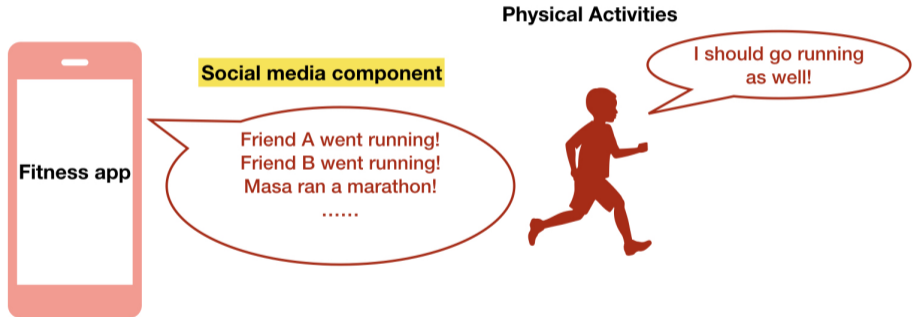
- ▶ Effect of motivational message on level of physical activities



- ▶ Seems to be easy: Run RCTs?

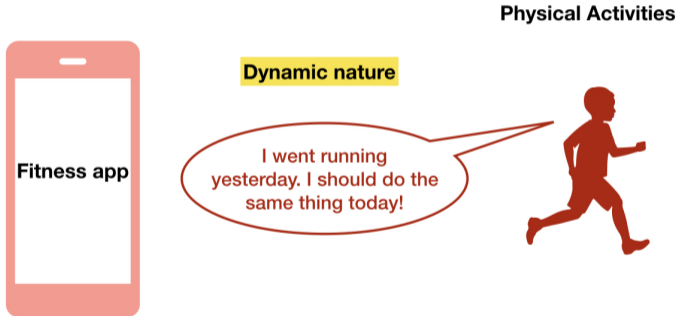
Interference

- ▶ An individual's outcome could depend on other individuals' outcomes.



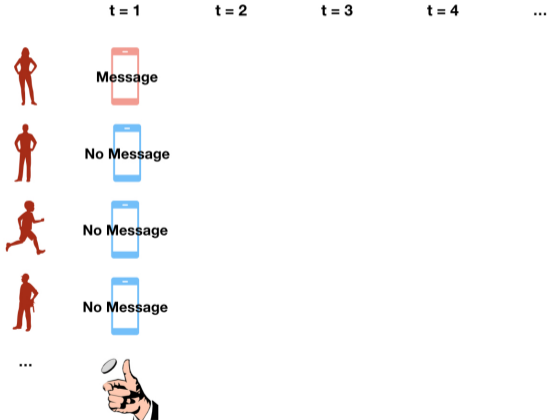
Dynamic Nature

- ▶ An additional time axis.



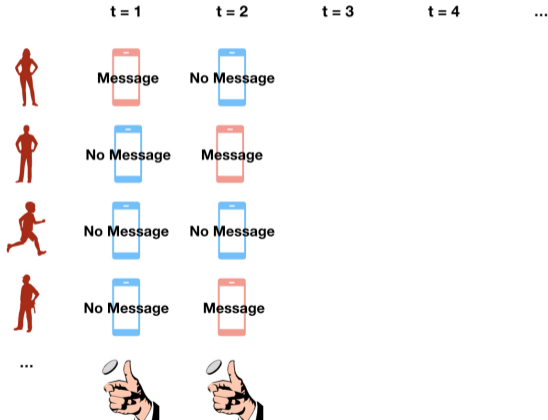
Micro-Randomized Trials

- ▶ The micro-randomized trial (MRT) is an experimental design that is often used to help evaluate and optimize dynamic interventions.
- ▶ We focus on Bernoulli treatments.



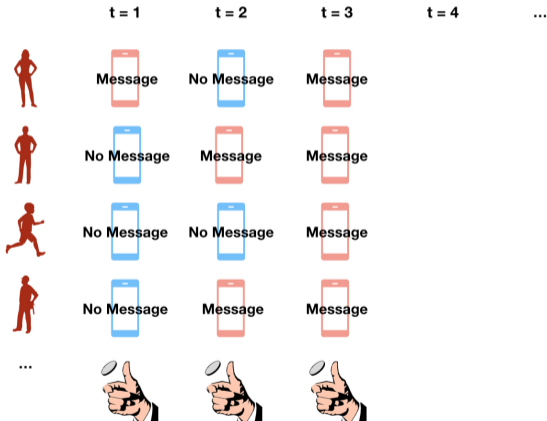
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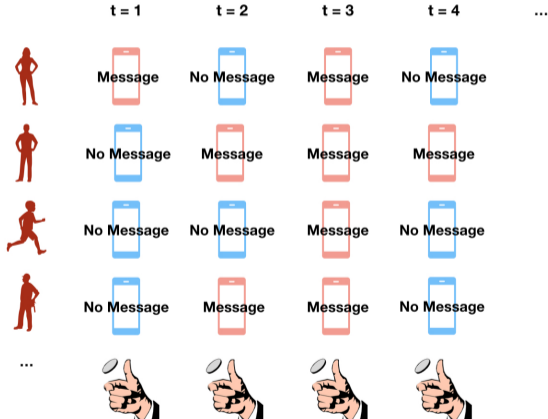
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Problem Setup

- ▶ There are n subjects of interests indexed by $i = 1, \dots, n$.
- ▶ Assume that there is an undirected graph with indices corresponding to the n subject. We call the undirected graph the *interference network* or the *interference graph*. We use $\{E_{ij}\}$ to denote the edge set of the graph.
- ▶ Let $\mathcal{N}_i = \{j : E_{ij} = 1\}$ be the set of neighbors of subject i .
- ▶ Let $Y_{it} \in \{0, 1\}$ denote the outcome of interest at time t .
- ▶ Let $W_{it} \in \{0, 1\}$ be the treatment at time t .

Assumptions

Assumption 1 (MDP with Network Interference)

Each unit $i = 1, \dots, n$ is characterized by an activation function $f_i(\cdot)$ such that, conditionally on $Y_{1:t}$ and $W_{1:t}$,

$$Y_{i(t+1)} \sim \text{Ber}(f_i(Y_{it}, W_{it}, Z_{it})) \text{ independently,}$$

where $Z_{it} = \sum_{j \in \mathcal{N}_i} Y_{jt}$, and $0 < f_i(y, w, z) < 1$ for all $y, w \in \{0, 1\}$ and $z \in \mathbb{R}_+$.

- ▶ Fitness app example: The probability of an individual goes running tomorrow depends on
 - whether she went running today (Y_{it}),
 - whether she receives any encouraging message from the app (W_{it}),
 - the total number of her friends that went running today (Z_{it}),
 - her individual characteristics (f_i).

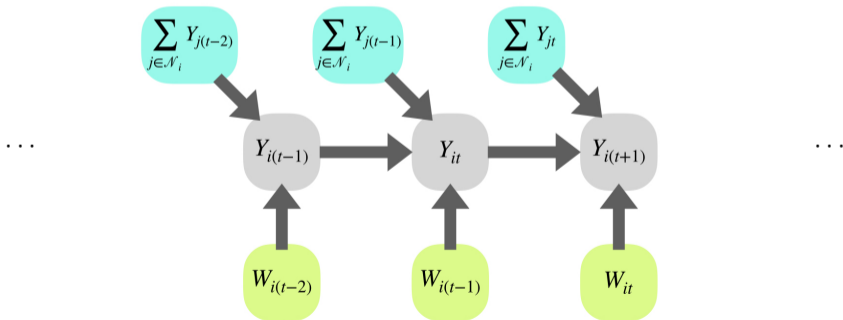
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Assumptions

Assumption 2 (Bernoulli treatment)

The treatments $W_{it} \sim \text{Ber}(\pi_i)$ independently for each i and each t .

Causal Estimands

- ▶ **Short-term direct effect.** The short-term direct effect quantifies the immediate effect of the unit's treatment on its own outcome.

$$\tau_{\text{SDE},t} = \frac{1}{n} \sum_{i=1}^n f_i(Y_{it}, 1, Z_{it}) - f_i(Y_{it}, 0, Z_{it}).$$

- ▶ **Long-term direct effect.** The long-term direct effect captures the long-term effect of the unit's treatment on its own outcome, averaged over units.

$$\tau_{\text{LDE}}(\gamma_1, \gamma_2) = \frac{1}{n} \sum_{i=1}^n (\mathbb{E}_{\mu(\pi_i=\gamma_1, \pi_{-i})} [Y_i] - \mathbb{E}_{\mu(\pi_i=\gamma_2, \pi_{-i})} [Y_i]).$$

Here $\mu(\pi_i = \gamma, \pi_{-i})$ stands for the stationary distribution of the MDP when the treatment probability of the i -th unit has been changed to γ .

Causal Estimands

► **Long-term total effect.**

It measures the effect of changing the entire treatment vector from π_2 to π_1 on the expected average outcome under the stationary distribution.

$$\tau_{\text{LTE}}(\pi_1, \pi_2) = \frac{1}{n} \sum_{i=1}^n (\mathbb{E}_{\mu(\pi_1)} [Y_i] - \mathbb{E}_{\mu(\pi_2)} [Y_i]).$$

Existence of a stationary distribution

Since both Y_{it} and W_{it} are binary random variables, the function f_i can be decomposed into four terms:

$$f_i(y, w, z) = a_i(z) + b_i(z)w + c_i(z)y + d_i(z)wy.$$

Additional assumptions:

- ▶ **Boundedness and Lipschitzness.** Each function f_i is Lipschitz with Lipschitz constant L_n in its third argument. The functions c_i and d_i satisfy $|c_i(z) + d_i(z)w| \leq B$ for any $z \in \mathbb{R}_+$ and any $w \in \{0, 1\}$.
- ▶ **Node Degree.** The largest node degree of the interference graph is bounded by D_n .
- ▶ **Contraction** The constants B , L_n and D_n satisfy $B + L_n D_n \leq C < 1$.

Example: Erdős-Rényi

Each edge is included in the interference graph with probability ρ_n , independently from every other edge, i.e., $E_{ij} \sim \text{Bernoulli}(\rho_n)$ independently.

$$L_n \sim \frac{1}{n\rho_n}, \quad D_n \sim n\rho_n.$$

Existence of a stationary distribution

Proposition (Stationary distribution)

1. Under the above Assumptions, there **exists** a stationary distribution $\mu(\pi)$, such that if $Y_0 \sim \mu(\pi)$, then $Y_t \sim \mu(\pi)$ for any $t \geq 0$.
2. Furthermore, the Markov chain (induced by the Bernoulli policy) is ergodic, i.e., the stationary distribution $\mu(\pi)$ is **unique**. For any initial distribution of Y_0 , we have that $Y_t \Rightarrow \mu(\pi)$ as $t \rightarrow \infty$.

Mean-Field Characterization

- ▶ Difficulties:
 - Under the stationary distribution μ , outcome Y_i 's are not independent.
 - The state space is of size 2^n .
- ▶ Consider a dynamical system—a “mean-field” version of the Markov Chain.
- ▶ Let $P_t = (P_{1t}, P_{2t}, \dots, P_{nt}) \in [0, 1]^n$ be the state of the dynamical system at time t . The evolution rule of the system is the following:

$$\begin{aligned} P_{i(t+1)} &= f_i(P_{it}, \pi_i, Q_{it}) \\ &= a_i(Q_{it}) + b_i(Q_{it}) \pi_i + c_i(Q_{it}) P_{it} + d_i(Q_{it}) \pi_i P_{it}, \end{aligned}$$

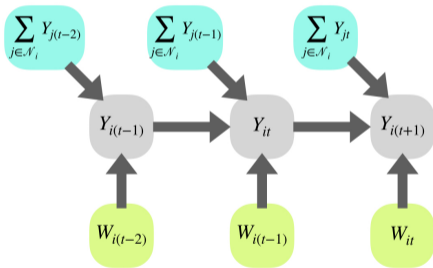
where $Q_{it} = \sum_{j \in \mathcal{N}_i} P_{jt}$.

- ▶ We can interpret $P_{i,t}$ as the **probability** that $Y_{i,t} = 1$ given past information. Here, the state of the i -th user depends directly on its neighbors probabilities rather than their realized outcomes.

Mean-Field Characterization

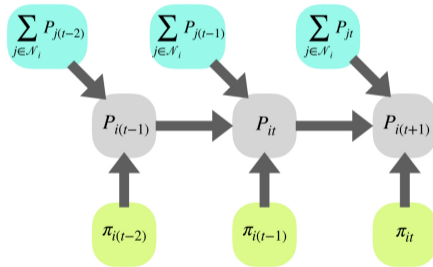
Markov chain

$$Y_{i(t+1)} \sim \text{Ber} \left(f_i \left(Y_{it}, W_{it}, \sum_{j \in \mathcal{N}_i} Y_{jt} \right) \right)$$



Dynamical system

$$P_{i(t+1)} = f_i \left(P_{it}, \pi_i, \sum_{j \in \mathcal{N}_i} P_{jt} \right)$$



Mean-Field Characterization

- ▶ Easier to study
 - The probabilities are non-random numbers, thus if the outcomes $Y_{it} \sim \text{Bern}(P_{it})$ independently, then the Y_{it} 's are independent.
 - The fixed point of the dynamical system can be characterized by a vector of length n .

Proposition (Fixed point)

1. There **exists** a fixed point $P^* \in [0, 1]^n$ of the dynamical system, i.e., if $P_t = P^*$, then $P_{t+1} = P^*$.
2. The fixed point is **unique**, and for any value of P_0 , we have $P_t \rightarrow P^*$ as $t \rightarrow \infty$.

How good is the mean-field approximation?

- ▶ $Y \sim \mu(\pi)$, where $\mu(\pi)$ is the stationary distribution of the Markov chain.
- ▶ $Y_i^* \sim \text{Ber}(P_i^*)$ independently, where P^* is the fixed point of the dynamical system.
- ▶ Is Y close to Y^* in distribution?
- ▶ Define an L_1 -Wasserstein distance between two laws ν_1 and ν_2

$$W_{L_1}(\nu_1, \nu_2) = \inf \{ \mathbb{E} [\|X_1 - X_2\|_1] : \mathcal{L}(X_1) = \nu_1, \mathcal{L}(X_2) = \nu_2 \}.$$

Theorem (Mean field approximation)

$$W_{L_1}(\mu, \mathcal{L}(Y^*)) \leq n\sqrt{L_n}\sqrt{C}/(2(1-C))$$

Short-term direct effect

- ▶ Recall that

$$\tau_{\text{SDE},t} = \frac{1}{n} \sum_{i=1}^n f_i(Y_{it}, 1, Z_{it}) - f_i(Y_{it}, 0, Z_{it}).$$

- ▶ A natural estimator to use here is the inverse propensity weighted (IPW) estimator:

$$\hat{\tau}_{\text{IPW},t} = \frac{1}{n} \sum_{i=1}^n Y_{i(t+1)} \left(\frac{W_t}{\pi_i} - \frac{1 - W_t}{1 - \pi_i} \right).$$

- ▶ The IPW estimator is consistent for the short-term direct effect.

Theorem (Short-term direct effect estimation)

$$\hat{\tau}_{\text{IPW}} = \tau_{\text{SDE},t} + \mathcal{O}_p\left(\frac{1}{\sqrt{n}}\right).$$

Long-term direct effect

- ▶ Recall that

$$\tau_{\text{LDE}}(\gamma_1, \gamma_2) = \frac{1}{n} \sum_{i=1}^n (\mathbb{E}_{\mu(\pi_i=\gamma_1, \pi_{-i})} [Y_i] - \mathbb{E}_{\mu(\pi_i=\gamma_2, \pi_{-i})} [Y_i]).$$

- ▶ Under this stationary distribution,

$$\begin{aligned} & \mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [Y_i] \\ &= \mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [f_i(Y_i, W_i, Z_i)] \\ &= \mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [a_i(Z_i) + b_i(Z_i)W_i + c_i(Z_i)Y_i + d_i(Z_i)W_iY_i] \\ &\approx \mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [a_i(Z_i)] + \mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [b_i(Z_i)] \gamma \\ &\quad + \mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [c_i(Z_i)] \mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [Y_i] + \mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [d_i(Z_i)] \gamma \mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [Y_i], \end{aligned}$$

if W_i , Y_i and Z_i are roughly independent.

Long-term direct effect

- ▶ Moving all terms involving $\mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [Y_i]$ to the left hand side, we get

$$\mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [Y_i] \approx \frac{\mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [a_i(Z_i)] + \mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [b_i(Z_i)] \gamma}{1 - \mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [c_i(Z_i)] - \mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [d_i(Z_i)] \gamma}.$$

- ▶ If we assume that Z_i 's are not influenced too much by the treatment probability of unit i , then

$$\mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [Y_i] \approx \frac{\mathbb{E}_{\mu(\pi)} [a_i(Z_i)] + \mathbb{E}_{\mu(\pi)} [b_i(Z_i)] \gamma}{1 - \mathbb{E}_{\mu(\pi)} [c_i(Z_i)] - \mathbb{E}_{\mu(\pi)} [d_i(Z_i)] \gamma}.$$

Long-term direct effect

The above heuristic in fact correctly recovers the long-term direct effect.

Theorem (Long-term direct effect characterization)

$$\mathbb{E}_{\mu(\pi_i=\gamma, \pi_{-i})} [Y_i] = \frac{\mathbb{E}_{\mu(\pi)} [a_i(Z_i) + b_i(Z_i)\gamma]}{\mathbb{E}_{\mu(\pi)} [1 - c_i(Z_i) - d_i(Z_i)\gamma]} + \mathcal{O}(\sqrt{L_n}).$$

- ▶ It suffices to estimate $\mathbb{E}_{\mu(\pi)} [a_i(Z_{it})], \dots, \mathbb{E}_{\mu(\pi)} [d_i(Z_{it})]$ for each i .

Long-term direct effect

► It suffices to estimate $\mathbb{E}_{\mu(\pi)} [a_i(Z_{it})], \dots, \mathbb{E}_{\mu(\pi)} [d_i(Z_{it})]$ for each i .

► Recall that

$$f_i(y, w, z) = a_i(z) + b_i(z)w + c_i(z)y + d_i(z)wy.$$

► Take $\mathbb{E}_{\mu(\pi)} [a_i(Z_{it})]$ as an example. Let

$$\hat{a}_i = \frac{\frac{1}{T} \sum_{t=1}^T Y_{i(t+1)}(1 - W_{it})(1 - Y_{it})}{\frac{1}{T} \sum_{t=1}^T (1 - W_{it})(1 - Y_{it})}.$$

Proposition

$$\mathbb{E} \left[(\hat{a}_i - \mathbb{E}_{\mu(\pi)} [a_i(Z_{it})])^2 \right] \leq C_1 \left(\frac{1}{T} + L_n \right),$$

for some constant C_1 not depending on i or n .

► Challenge: Y_{it} 's are not independent.

Long-term direct effect

► Let

$$\hat{\tau}_{\text{LDE}}(\gamma_1, \gamma_2) = \frac{1}{n} \sum_{i=1}^n \frac{\hat{a}_i + \hat{b}_i \gamma_1}{1 - \hat{c}_i - \hat{d}_i \gamma_1} - \frac{1}{n} \sum_{i=1}^n \frac{\hat{a}_i + \hat{b}_i \gamma_2}{1 - \hat{c}_i - \hat{d}_i \gamma_2}.$$

Corollary (Long-term direct effect estimation)

$$\hat{\tau}_{\text{LDE}}(\gamma_1, \gamma_2) - \tau_{\text{LDE}}(\gamma_1, \gamma_2) = \mathcal{O}_p \left(\frac{1}{\sqrt{T}} + \sqrt{L_n} \right).$$

Long-term total effect

- ▶ Recall that

$$\tau_{\text{LTE}}(\pi + \Delta \mathbf{v}, \pi) = \frac{1}{n} \sum_{i=1}^n (\mathbb{E}_{\mu(\pi + \Delta \mathbf{v})} [Y_i] - \mathbb{E}_{\mu(\pi)} [Y_i]).$$

where $\|\mathbf{v}\| = \sqrt{n}$.

- ▶ Hard to analyze the stationary distribution: start with fixed point of the dynamical system
- ▶ When Δ is small,

$$P_i^*(\pi + \Delta \mathbf{v}) - P_i^*(\pi) \approx \Delta \frac{d}{d\Delta} P_i^*(\pi + \Delta \mathbf{v}) \approx \Delta (\nabla_{\pi} P_i^*(\pi)^{\top} \mathbf{v}).$$

- ▶ Recall that P^* satisfies

$$P_i^*(\pi) = a_i (Q_i^*(\pi)) + b_i (Q_i^*(\pi)) \pi_i + c_i (Q_i^*(\pi)) P_i^*(\pi) + d_i (Q_i^*(\pi)) \pi_i P_i^*(\pi),$$

where $Q_i^*(\pi) = \sum_{j \in \mathcal{N}_i} P_j^*(\pi)$.

Long-term total effect

- ▶ $P_i^*(\pi) = a_i(Q_i^*(\pi)) + b_i(Q_i^*(\pi))\pi_i + c_i(Q_i^*(\pi))P_i^*(\pi) + d_i(Q_i^*(\pi))\pi_i P_i^*(\pi)$,
- ▶ Let $p_i^*(\pi) = \nabla_{\pi} P_i^*(\pi)^T \mathbf{v}$. Taking derivative on both hand side gives:

$$p_i^*(\pi) = b_i(Q_i^*(\pi))v_i + d_i(Q_i^*(\pi))P_i^*(\pi)v_i + [c_i(Q_i^*(\pi)) + d_i(Q_i^*(\pi))\pi_i] p_i^*(\pi) + [a'_i(Q_i^*(\pi)) + b'_i(Q_i^*(\pi))\pi_i + c'_i(Q_i^*(\pi))P_i^*(\pi) + d'_i(Q_i^*(\pi))\pi_i P_i^*(\pi)] \sum_{j \in \mathcal{N}_i} p_j^*(\pi).$$

- ▶ We have a set of linear equations for $p_i^*(\pi)$'s.
- ▶ Solving for $p_i^*(\pi)$ gives

$$\mathbf{p}^*(\pi) = (\mathbf{I} - \mathbf{DA} - \mathbf{W})^{-1} \mathbf{u},$$

where A is the adjacency matrix of the interference graph,

$$\mathbf{D} = \text{diag}(a'_i(Q_i^*(\pi)) + b'_i(Q_i^*(\pi))\pi_i + c'_i(Q_i^*(\pi))P_i^*(\pi) + d'_i(Q_i^*(\pi))\pi_i P_i^*(\pi)),$$

$$\mathbf{W} = \text{diag}(c_i(Q_i^*(\pi)) + d_i(Q_i^*(\pi))\pi_i), \text{ and } \mathbf{u} = \text{vec}(v_i(b_i(Q_i^*(\pi)) + d_i(Q_i^*(\pi))P_i^*(\pi))).$$

Long-term total effect

- ▶ Summarizing the above findings, we have

$$\begin{aligned}\tilde{\tau}_{\text{LTE}}(\pi + \Delta \mathbf{v}, \pi) &= \frac{1}{n} \sum_{i=1}^n (P_i^*(\pi + \Delta \mathbf{v}) - P_i^*(\pi)) \approx \frac{\Delta}{n} \sum_{i=1}^n (\nabla_{\pi} P_i^*(\pi)^{\top} \mathbf{v}) = \frac{\Delta}{n} \sum_{i=1}^n p_i^*(\pi) \\ &= \frac{\Delta}{n} \mathbf{1}^{\top} \mathbf{p}^*(\pi) = \frac{\Delta}{n} \mathbf{1}^{\top} (I - DA - W)^{-1} \mathbf{u}.\end{aligned}$$

- ▶ Questions remaining:
 - Is $\tau_{\text{LTE}}(\pi + \Delta \mathbf{v}, \pi)$ close to $\tilde{\tau}_{\text{LTE}}(\pi + \Delta \mathbf{v}, \pi)$?
 - How good is the above approximation?

Long-term total effect

Theorem

$$\tau_{\text{LTE}}(\pi + \Delta \mathbf{v}, \pi) = \tilde{\tau}_{\text{LTE}}(\pi + \Delta \mathbf{v}, \pi) + \mathcal{O}\left(\sqrt{L_n}\right).$$

Theorem (Long-term total effect characterization)

Under some additional smoothness assumption,

$$\frac{\tau_{\text{LTE}}(\pi + \Delta_n \mathbf{v}, \pi)}{\Delta_n} = \frac{1}{n} \mathbf{1}^\top (I - DA - W)^{-1} \mathbf{u} + \mathcal{O}\left(\frac{\sqrt{L_n}}{\Delta_n} + \Delta_n\right).$$

Long-term total effect

- ▶ The above shows that

$$\tau_{\text{LTE}}(\pi + \Delta \mathbf{v}, \pi) \approx \frac{\Delta}{n} \mathbf{1}^\top (I - DA - W)^{-1} \mathbf{u}.$$

- ▶ The estimation strategy is to estimate the components D , W and \mathbf{u} separately.
- ▶ This involves estimating $a'_i(Q_i^*(\pi))$, $b'_i(Q_i^*(\pi))$, \dots .
- ▶ We estimate $a'_i(Q_i^*(\pi))$ by running a regression of $Y_{i(t+1)}$ on Z_{it} conditioning on $Y_{it} = 0$ and $W_{it} = 0$ using data from different time points.
- ▶ Challenges:
 - a_i is not linear,
 - data points are not i.i.d., and thus standard results from linear regression do not apply directly.

Long-term total effect

- We define

$$\hat{\tau}_{\text{LTE}}(\pi + \Delta \mathbf{v}, \pi) = \frac{\Delta}{n} \mathbf{1}^\top (I - \hat{D}A - \hat{W})^{-1} \hat{\mathbf{u}}.$$

Theorem (Long-term total effect estimation)

Assume that Δ varies with n . We write Δ_n to emphasize such dependency.

$$\frac{\hat{\tau}_{\text{LTE}}(\pi + \Delta_n \mathbf{v}, \pi)}{\Delta_n} = \frac{\tau_{\text{LTE}}(\pi + \Delta_n \mathbf{v}, \pi)}{\Delta_n} + \tilde{O}_p \left(\frac{1}{\Delta_n \sqrt{D_n}} + \Delta_n + \left(\frac{D_n}{\sqrt{T}} + \frac{1}{\sqrt{D_n}} \right) \right).$$

Thank you!