

Experimenting under Stochastic Congestion

Shuangning Li

Harvard University



Joint work with



Ramesh Johari



Stefan Wager



Kuang Xu

Li, Shuangning, Ramesh Johari, Stefan Wager, and Kuang Xu. **Experimenting under Stochastic Congestion.** *arXiv preprint arXiv:2302.12093* (2023).

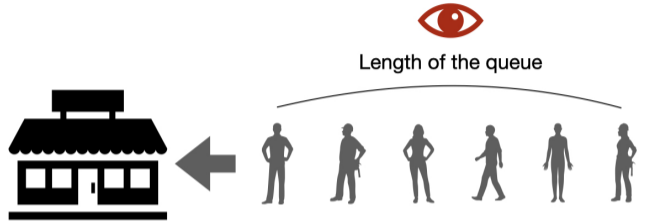
Toy example: ice cream shop



Price p

Customers wait in line to buy the ice cream

Toy example: ice cream shop



Price p



Customers wait in line to buy the ice cream

New Customer

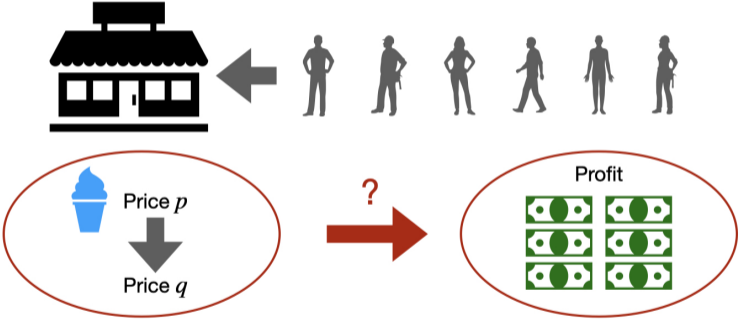


The queue is too long: I'll wait too long to get my ice cream.

The price is too high.

This store seems popular!

As the shop owner...



As the shop owner...



More customers are attracted to the store.

The line gets longer.

This further impacts customer behaviors.

The key question

How does ρ impact the (expected) number of arrivals? (1)

The key question

How does p impact the (expected) number of arrivals? (1)

- ▶ We direct our attention to estimating the **derivative** of the (expected) number of arrivals to the system with respect to the system parameter p .
- ▶ Why derivative?
 - Small perturbation to the system: protect from risk.
 - Can perform (stochastic) gradient descent to optimize the system.

Two other examples

► **Fast-track in the emergency department.**

To address the problem of long waits in the emergency department, a hospital implements a fast-track system. Patients with minor illness can go to the “fast-track” area and be treated faster.

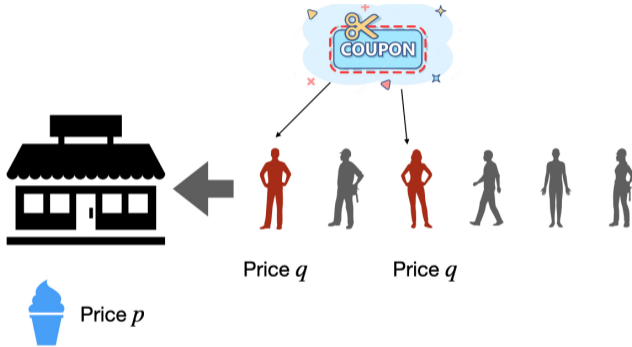
- The parameter p can be one of various quality measures of the fast-track service (e.g., quality of equipment and doctors, efficiency of the service).
- Question (1) can help the hospital understand how the quality of the fast-track service affects congestion in the emergency department.

Two other examples

- ▶ **Flight check-in.** Passengers checking in for their flights have a few options: wait in line to check in in person, check in at the self-service kiosks, or check in online.
 - In order to reduce congestion, the airline sends emails and posts signs encouraging passengers to use online or kiosk check-in. We let the parameter p measure the level of this encouragement.
 - Question (1) studies how much the encouragement actions impact the number of in-person check-ins and alleviate congestion at the airport.

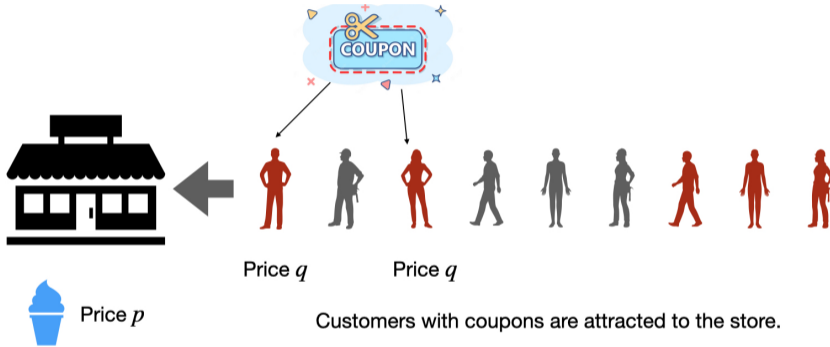
Is this problem trivial?

Can we run an A/B test?



Is this problem trivial?

Can we run an A/B test?



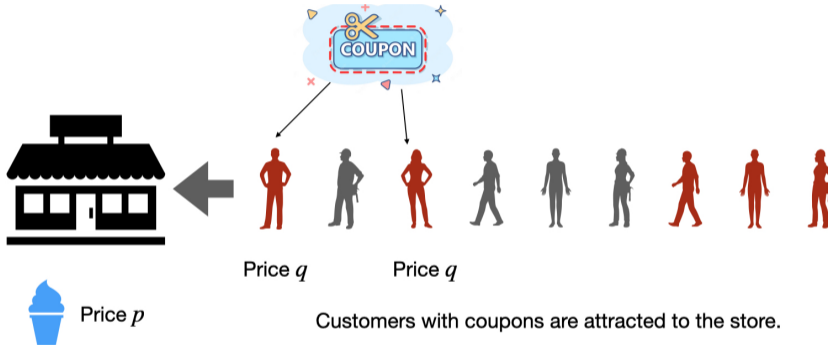
Customers with coupons are attracted to the store.

The line gets longer.

Impacts other customers.

Is this problem trivial?

Can we run an A/B test?



Price q

Price q

Price p

Customers with coupons are attracted to the store.

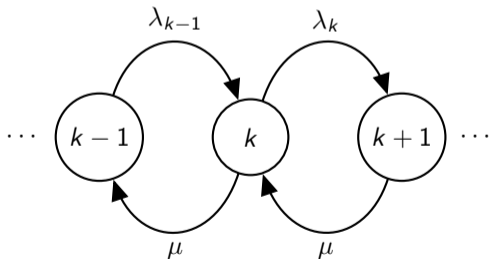
The line gets longer.

Impacts other customers.

Interference exists!

Setup

- ▶ A single-server queueing system that runs in continuous time.
- ▶ Service rate μ .
- ▶ Arrival rate $\lambda_k(p)$, where k is the current queue length.
- ▶ To understand $\lambda_k(p)$, we can think of the system as consisting of a queue with an outside option. We can take $\lambda_k(p) = \lambda_{\text{raw}} \text{prob}(k, \mu, p)$

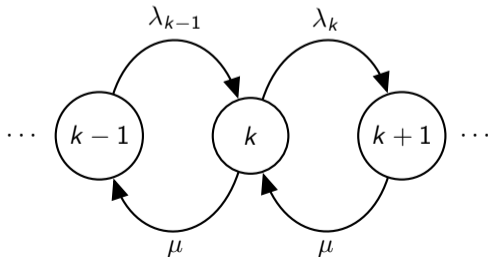


Setup

- ▶ The queue length process can be thought of as a continuous-time Markov chain.
- ▶ Let $\pi_k(p)$ denote the steady-state probability of being in state k .
- ▶ Let $\bar{\lambda}(p)$ be the average arrival rate in steady-state:


$$\bar{\lambda}(p) = \sum_{k=0}^{\infty} \pi_k(p) \lambda_k(p).$$

- ▶ We are interested in the derivative $d\bar{\lambda}/dp$.
- ▶ We call $d\bar{\lambda}/dp$ the *estimand* of interest.



Setup

$$\frac{d}{dp} \bar{\lambda}(p) = \frac{d}{dp} \left(\sum_{k=0}^{\infty} \pi_k(p) \lambda_k(p) \right) = \sum_{k=0}^{\infty} \pi'_k(p) \lambda_k(p) + \sum_{k=0}^{\infty} \pi_k(p) \lambda'_k(p)$$



Indirect effect Direct effect

- ▶ A/B tests can only be used to estimate $\lambda'_k(p)$.
- ▶ Can we write $\frac{d}{dp} \bar{\lambda}(p)$ in terms of $\lambda'_k(p)$ only?

Three representations of the estimand

Three representations of the estimand

$$\frac{d\bar{\lambda}(p)}{dp} = -\mu\pi'_0(p) = \mu\pi_0(p) \sum_{k=0}^{K-1} \frac{\lambda'_k(p)}{\lambda_k(p)} \sum_{i=k+1}^K \pi_i(p).$$

We need exponential interarrival times.

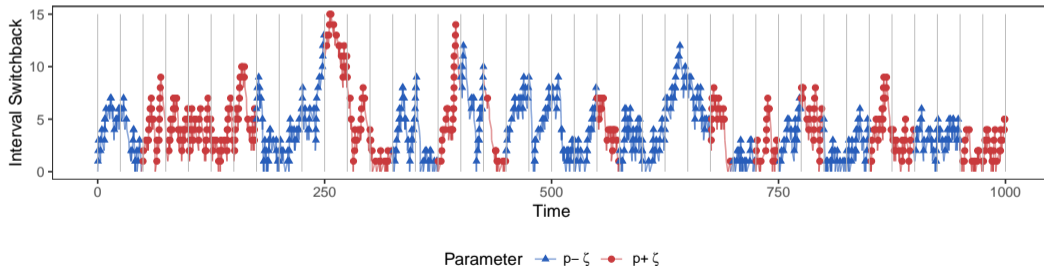
This holds under more general assumptions.

Finite difference

- ▶ The goal of estimation is a derivative with respect to the parameter p .
- ▶ Use finite differences.
- ▶ Specifically, let $\zeta > 0$ be a small perturbation of the parameter p . We compute finite differences by contrasting the behavior of the queue when the parameter is set to be $p + \zeta$ and $p - \zeta$.

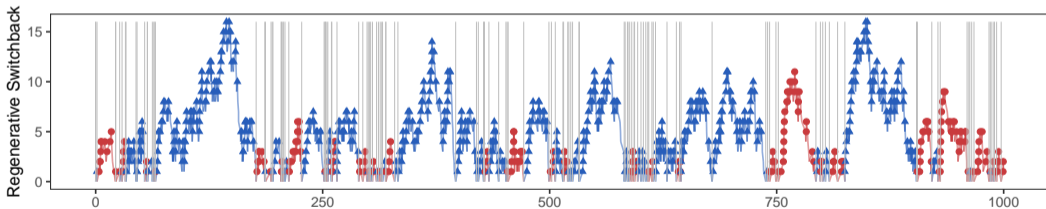
Experimental design

- ▶ Interval switchback experiment.
- ▶ We divide the time horizon into blocks of equal length l .
- ▶ At each time $t = zl$ for some $z \in \mathbb{Z}_+$, the parameter is set to be $p + \zeta$ or $p - \zeta$ with probability $1/2$ independently, and the parameter is kept the same in the time block $[zl, (z + 1)l)$ [Bojinov et al., 2022, Hu and Wager, 2022].



Experimental design

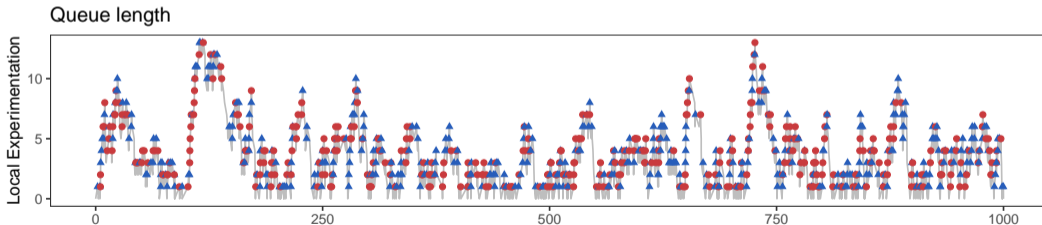
- ▶ Regenerative switchback experiment.
- ▶ The parameter can only be changed when the queue length is k_r .
- ▶ At each visit of state k_r , the parameter is set be $p + \zeta$ or $p - \zeta$ with probability 1/2 independently [Johari, 2019, Glynn et al., 2022].



Experimental design

- ▶ Local perturbation experiment.
- ▶ Instead of changing the parameter for all customers at the same time, we randomize the parameter for each customer. Specifically, let the parameter for customer i be

$$p_i = p + \zeta \varepsilon_i, \quad \text{where } \varepsilon_i \stackrel{\text{iid}}{\sim} \{\pm 1\} \text{ uniformly at random.}$$



Estimators

Three representations of the estimand

$$\frac{d\bar{\lambda}(p)}{dp} = -\mu\pi'_0(p) = \mu\pi_0(p) \sum_{k=0}^{K-1} \frac{\lambda'_k(p)}{\lambda_k(p)} \sum_{i=k+1}^K \pi_i(p).$$

Target representation of estimand	$\frac{d}{dp} \bar{\lambda}(p)$	$-\mu\pi'_0(p)$	$\mu\pi_0(p) \sum_{k=0}^{K-1} \frac{\lambda'_k(p)}{\lambda_k(p)} \sum_{i=k+1}^K \pi_i(p)$
Interval switchback	$\hat{\tau}_{IS, \bar{\lambda}}$	$\hat{\tau}_{IS, \pi_0}$	—
Regenerative switchback	$\hat{\tau}_{RS, \bar{\lambda}}$	$\hat{\tau}_{RS, \pi_0}$	—
Local perturbation	—	—	$\hat{\tau}_{LE}$

Asymptotic Behaviors of the Estimators

Central Limit Theorems

For any method a ,

$$\sqrt{T\zeta_T^2} \left(\hat{\tau}_a(T, \zeta_T) - \frac{d}{dp} \bar{\lambda}(p) \right) \Rightarrow \mathcal{N}(0, \sigma_a^2(p)).$$

- ▶ We give explicit formulas for each $\sigma_a^2(p)$.
- ▶ We provide variance estimators and get confidence intervals.

Comparison of the estimators

Variance comparison

$$\sigma_{\text{LE}}^2(\rho) = \frac{1}{2} \sigma_{\pi_0}^2(\rho),$$

and

$$\sigma_{\text{LE}}^2(\rho) \leq \sigma_{\bar{\lambda}}^2(\rho),$$

The equality holds if and only if the queue is an $M/M/1$ queue when the parameter is set at ρ , i.e., $\lambda_k(\rho) = \lambda_0(\rho)$ for all $k \geq 0$.

Target representation of estimand	$\frac{d}{d\rho} \bar{\lambda}(\rho)$	$-\mu \pi'_0(\rho)$	$\mu \pi_0(\rho) \sum_{k=0}^{K-1} \frac{\lambda'_k(\rho)}{\lambda_k(\rho)} \sum_{i=k+1}^K \pi_i(\rho)$
Interval switchback	$\hat{\tau}_{\text{IS}, \bar{\lambda}}$	$\hat{\tau}_{\text{IS}, \pi_0}$	—
Regenerative switchback	$\hat{\tau}_{\text{RS}, \bar{\lambda}}$	$\hat{\tau}_{\text{RS}, \pi_0}$	—
Local perturbation	—	—	$\hat{\tau}_{\text{LE}}$

This is better!

Why is $\hat{\tau}_{LE}$ better?

Exploiting the structure of the queueing model

Target representation of estimand	$\frac{d}{dp} \bar{\lambda}(p)$	$-\mu \pi'_0(p)$	$\mu \pi_0(p) \sum_{k=0}^{K-1} \frac{\lambda'_k(p)}{\lambda_k(p)} \sum_{i=k+1}^K \pi_i(p)$
Interval switchback	$\hat{\tau}_{IS, \bar{\lambda}}$	$\hat{\tau}_{IS, \pi_0}$	—
Regenerative switchback	$\hat{\tau}_{RS, \bar{\lambda}}$	$\hat{\tau}_{RS, \pi_0}$	—
Local perturbation	—	—	$\hat{\tau}_{LE}$

Conditioning on the state of the queue

Thank you!

